

Solution to Assignment 11

9.42: **(a)** Under  $H_0$ ,  $p_i(p) = P(X = i) = \binom{5}{i} p^i (1-p)^{5-i}$ . So the likelihood for our multinomial data  $(Y_0, \dots, Y_5)$  under  $H_0$  is

$$L(p) = \binom{280}{157, \dots, 1} \prod_{i=0}^5 \left[ \binom{5}{i} p^i (1-p)^{5-i} \right] = \binom{280}{157, \dots, 1} 2500 p^{199} (1-p)^{1201}.$$

Solving

$$l'(\hat{p}) = \frac{199}{\hat{p}} - \frac{1201}{1-\hat{p}} = 0,$$

$\hat{p} = 199/1400$  is the maximum likelihood estimate for  $p$ . **(b)** With the pooling,  $O_0 = 157, \dots, O_2 = 35$ , and  $O_3 = 17+1+1 = 19$ . Then  $E_i = 280 p_i(\hat{p})$  for  $i = 0, 1, 2$ , and  $E_3 = 280 - E_0 - E_1 - E_2$ . Pearson's chi-square is

$$\sum_{i=0}^3 \frac{(O_i - E_i)^2}{E_i} = 44.148.$$

The upper 5th percentile for the chi-square distribution on 2 degrees of freedom is 5.99, so we can easily reject the hypothesis that distribution for the data is binomial. **(c)** Let  $Y_i$  denote the  $i$ -th variable (since we are using  $X$ 's for the multinomial counts). Then

$$-2 \log \Lambda = \sum_{i=1}^{280} \frac{(Y_i - 5\hat{p})^2}{5\hat{p}(1-\hat{p})} = \frac{\sum Y_i^2 - 10\hat{p} \sum Y_i + 280\hat{p}^2}{5\hat{p}(1-\hat{p})} = 429.125$$

This can be compared against the upper 5-th percentile for the chi-square distribution on 279 degrees of freedom (although the approximation seems suspect here). Since this distribution has mean 279 and variance 558, by normal approximation this quantile is approximately  $279 + 1.645\sqrt{558} = 317.86$ , and we would reject  $H_0$ .

11.24: The mass function for  $U_Y$  under  $H_0$  is

u	0	1	2	3	4	5	6
$P_{H_0}(U_Y = u)$	0.1	0.1	0.2	0.2	0.2	0.1	0.1

11:40: **(b)**  $\bar{X} = 17.23, \bar{Y} = 25.00, s_x^2 = 32.16, s_y^2 = 10.08,$

$$s_p^2 = \frac{9s_x^2 + 9s_y^2}{18} = 21.12 = (4.595)^2,$$

$$s_{\bar{X}-\bar{Y}} = s_p \sqrt{\frac{1}{10} + \frac{1}{10}} = 2.055,$$

and  $t_{18}(0.025) = 2.101$ . So the confidence interval is

$$(\bar{X} - \bar{Y} \pm t_{18}(0.025)s_{\bar{X}-\bar{Y}}) = (-7.77 \pm 4.32).$$

(c) The  $t$ -statistic to test  $H_0$  is  $(\bar{X} - \bar{Y})/s_{\bar{X}-\bar{Y}} = -3.78$ . Since the magnitude of this statistic exceeds  $t_{18}(0.025) = 2.101$ ,  $H_0$  is rejected. (d) By direct calculation,  $T = 67$  and so

$$U = T - \frac{m(m+1)}{2} = 12.$$

Under  $H_0$ ,  $E(U) = mn/2 = 50$  and  $\text{Var}(U) = mn(m+n+1)/12 = 175$ . So

$$Z = \frac{U - 50}{\sqrt{175}} = -2.87.$$

Since the magnitude of  $Z$  exceeds  $z(0.025) = 1.96$ , the nonparametric test would also reject  $H_0$ .

14.2:  $\bar{x} = -0.046$ ,  $\bar{y} = -0.075$ ,

$$\sum (x_i - \bar{x})^2 = 10.41304, \quad \sum (y_i - \bar{y})^2 = 8.92665,$$

and

$$\sum (x_i - \bar{x})(y_i - \bar{y}) = 9.4177.$$

So

$$b = \frac{9.4177}{10.41304} = 0.9044, \quad a = -0.075 + (0.9044)(0.046) = -0.0334$$

and

$$d = \frac{9.4177}{8.92665} = 1.0550 \quad c = -0.046 + (1.0550)(0.075) = 0.0331.$$

The two lines are different. They would only be the same if the data were perfectly correlated.